Missing Responses in Longitudinal Data

Example: Study on the Financial Crisis

The data that we analyzed...

id	Age	Sex	Y2	Y3	Y4	Y5	Y6	Y7	Y8
10	23.92	F	2	2	3	2	3	3	3
17	76.09	Μ	2	4	2	4	4	4	4
29	54.04	F	1	2	2	4	3	3	1
÷	÷	÷	÷	÷	÷	÷	÷	÷	÷
1810	46.77	М	2	2	2	2	2	2	2
1826	24.77	М	2	2	2	2	2	3	1
1837	26.29	F	1	1	1	1	1	1	1

Example: Study on the Financial Crisis

The data that were collected...

Age	Sex	Y2	Y3	Y4	Y5	Y6	Y7	Y8
40.55	F	2	1	3	2	NA	3	NA
23.92	F	2	2	3	2	3	3	3
44.15	Μ	NA	NA	NA	NA	1	3	NA
÷	÷	÷	÷	÷	÷	÷	÷	÷
29.58	Μ	4	4	NA	NA	2	4	NA
54.64	Μ	2	1	1	1	1	1	NA
30	Μ	2	1	1	NA	NA	NA	NA
	Age 40.55 23.92 44.15 : 29.58 54.64 30	Age       Sex         40.55       F         23.92       F         44.15       M         29.58       M         54.64       M         30       M	AgeSexY240.55F223.92F244.15MNA29.58M454.64M230M2	AgeSexY2Y340.55F2123.92F2244.15MNANA29.58M4454.64M2130M21	AgeSexY2Y3Y440.55F21323.92F22344.15MNANANA29.58M44NA54.64M21130M211	AgeSexY2Y3Y4Y540.55F213223.92F223244.15MNANANANA29.58M44NANA54.64M21130M211NA	AgeSexY2Y3Y4Y5Y640.55F2132NA23.92F2232344.15MNANANANA129.58M44NANA254.64M21111	AgeSexY2Y3Y4Y5Y6Y740.55F2132NA323.92F22323344.15MNANANANA1329.58M4ANANA2454.64M2111130M21NANANANA

**Missing data** refer to any observations which we *intended* to collect, but which were not recorded in our data file for **any reason**.

**Missing data** is a pervasive problem across most domains, but it is *particularly* common in longitudinal studies.

# Classification of Missing Data Mechanisms

Notation for Missingness

We define an **observation indicator**  $R_{ij}$ .

$${\sf R}_{ij} = egin{cases} 1 & Y_{ij} ext{ is observed}, \ 0 & ext{otherwise}. \end{cases}$$

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We partition the **outcome**  $Y_i$  into the **observed components**  $Y_i^O$  and the **missing components**,  $Y_i^M$ .

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$$f_{R_i}(r_i|Y_i^{O}, Y_i^{M}, X_i) = f_{R_i}(r_i|X_i).$$

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- Otherwise, data are said to be not missing at random (NMAR).
  - Data may be NMAR if, for instance, individuals who smoke more (Y<sub>ij</sub> large) are less likely to continue responding to a smoking questionnaire.

# Patterns of Missingness



Impacts of Missingness

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- 3. Likelihood based techniques will be valid if the data are MAR or MCAR, so long as the model is correctly specified.
- 4. In all other situations, estimators will be **biased**, and inference will be **invalid**.

General Techniques for Handling Missingness

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- 2. Available data analysis, where all observations that were made are included in the data frame.
- 3. Weighting techniques, where pseudo datasets are created based on weighting the available information.
- 4. **Imputation techniques**, where the missing values are filled-in based on an underlying model.

# A Note on Handling NMAR Missingness

# The four classes of techniques listed above will **not** be valid for NMAR data. NMAR data **need** joint modelling strategies for

 $f(Y_i, R_i).$ 

Weighting Techniques

One special type of missingness is **dropout**. In this case, an individual is observed only until  $t_j$ , and no times after.

Define  $D_i$  to be the **dropout time** 

$$D_i = 1 + \sum_{j=1}^{K} R_{ij}.$$

#### Probability of Inclusion

We can think about estimating the **probability of inclusion** for any individual, at any time in the study.

 $\pi_{ij} = P(D_i > j | D_i \ge j, i).$ 

This gives the probability that individual i was still under observation at time j, assuming that they made it to at least time j - 1. We can estimate these probabilities via (e.g.) logistic regression.

Individuals with low  $\pi_{ij}$  were unlikely to have been observed. Individuals with high  $\pi_{ij}$  were likely to have been observed.

#### Balancing the Observed Data

**True Population** 

**Observed Data** 



# Constructing a Pseudo-Population

$$w_B = \frac{1}{\pi_B} = 2$$
 and  $w_M = \frac{1}{\pi_M} = 20.$ 

**Observed Data** 

**Pseudo Population** 





# Applying this to Longitudinal Data

We can take

$$\pi_i = P(D_i > K) = \prod_{j=1}^K \pi_{ij},$$

and correspondingly get

$$w_i = \frac{1}{\pi_i} = \frac{1}{\prod_{j=1}^K \pi_{ij}}.$$

Then, by running a **complete case analysis** with these weights, any MCAR or MAR missingness will be accounted for (assuming the  $\pi_i$  are correct).

#### More Efficiently

We can extend this idea to weighting **all available data** instead of just the complete cases.

$$w_{ij}=rac{1}{P(D_i>j)}=\left[\prod_{\ell=1}^j\pi_{i\ell}
ight]^{-1}.$$

Then giving each individual's observations at time j a weight of  $w_{ij}$ , and running an **available data analysis** produces more efficient corrected estimators.

# **IPW-GEE**

Applied specifically to GEE:

$$\sum_{i=1}^{N} D'_i V_i^{-1} W_i \{ Y_i - \mu_i(\beta) \} = 0.$$

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- We either need  $X_i$  fully observed or  $V_i$  diagonal.
  - Recall that V<sub>i</sub> need not be correctly specified. If X<sub>i</sub> is not fully observed, take V<sub>i</sub> diagonal.

Imputation Techniques

Using some model, estimate the missing values Y<sup>M</sup><sub>i</sub> based on the observed values, Y<sup>O</sup><sub>i</sub> and variates X<sub>i</sub>.

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#### Imputation in General

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Need to choose **single** or **multiple** imputation, and the **imputation procedure**.

#### Multiple Imputation

Repeat the imputation process *m* times, giving  $\widehat{\beta}^{(k)}$  for  $k = 1, \ldots, m$ . Then

$$\widehat{\beta} = \frac{1}{m} \sum_{k=1}^{m} \widehat{\beta}^{(k)},$$

and we further take

$$\widehat{\operatorname{cov}}\left(\widehat{\beta}\right) = \frac{1}{m} \sum_{k=1}^{m} \operatorname{cov}\left(\widehat{\beta}^{(k)}\right) + \frac{m+1}{m(m-1)} \sum_{k=1}^{m} \left(\widehat{\beta}^{(k)} - \widehat{\beta}\right) \left(\widehat{\beta}^{(k)} - \widehat{\beta}\right)'.$$

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- 5. Continue imputing, estimate  $\hat{\beta}^{(1)}$ , and then repeat *m* times.

#### Problem with Underestimating Uncertainty

This procedure outlined **underestimates** the variability that should be inherent to this imputation procedure, since  $\hat{Y}_{ij}$  is estimated not fixed!

Instead of using estimated  $\hat{\gamma}_j$ , we draw from the posterior distribution, giving  $\tilde{\gamma}_j$ , and otherwise proceed as outlined.

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- 4. Repeat this process for all j, and then m times.

# Any technique that uses **maximum likelihood** (e.g., GLMEMs or transition models) will result in valid inference if the data are **MCAR** or **MAR**. In this case

$$f(Y_i|X_i) = f(Y_i^{\mathsf{O}}|X_i) = f(Y_i^{\mathsf{M}}|X_i).$$

There are procedures (using Expectation Maximization (EM)) which make the connection between **likelihood** and **imputation** more clear.

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- Weighting techniques generate pseudo-populations that match the would-be observed population using estimated probabilities.
- Imputation techniques fill in the missing values based on specific regression models.