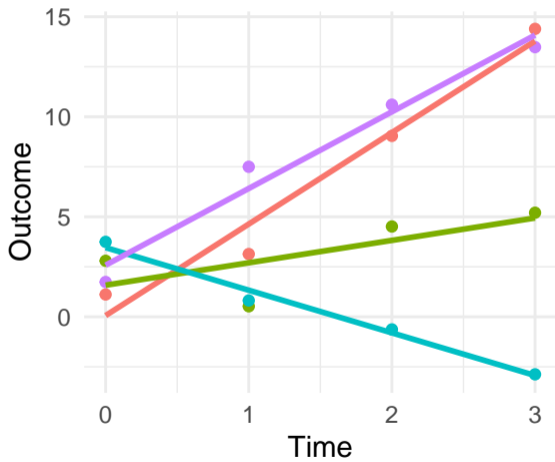


Linear Mixed Effects Models

What if we fit a *different* model for each individual?

We could find a unique β_{0i} and β_{1i} for each individual ...



Problems with this...

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2. This procedure would ignore **within-subject correlations**.

Alternative idea...

Share parameters.

Use **individual-level** terms in addition to **population-level** terms which are shared across the population.

Parameter Sharing in Practice

Instead of β_{0i} and β_{1i} , we could break these down into $\beta_{0i} = \beta_0 + b_{0i}$ and $\beta_{1i} = \beta_1 + b_{1i}$.

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This way,

$$\begin{aligned} Y_{ij} &= \beta_{0i} + \beta_{1i}t_{ij} + \epsilon_{ij} \\ &= (\beta_0 + b_{0i}) + (\beta_1 + b_{1i})t_{ij} + \epsilon_{ij} \\ &= \underbrace{(\beta_0 + \beta_1 t_{ij})}_{\text{Population Level}} + \underbrace{(b_{0i} + b_{1i} t_{ij})}_{\text{Individual Level}} + \epsilon_{ij}. \end{aligned}$$

Random Effects Terms

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 - ▶ The measurement **variation** at time j : ϵ_{ij} .

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Regression Parameters with a Distribution

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Can we do the same thing **here**?

Specification of a Linear Mixed Effects Model

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Typically, we will set $G_i = \sigma^2 I$ to:

1. maintain the interpretation as **sampling error**; and
2. ensure **identifiability**.

Mean, Variance, and Distribution

- ▶ If we condition on the **random effects** $E[Y_i|b_i] = X_i\beta + Z_ib_i$ and $\text{var}(Y_i|b_i) = \text{var}(\epsilon_i|b_i) = G_i$.

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- ▶ This is a *specific* form of a **linear marginal model**!

Specific Examples: Random Intercept and Random Slope Models

The Random Intercept Model

The most basic version of a mixed effects model takes $Z_{ij} = 1$, and as such, b_j is a scalar for each individual.

This is called the **random intercept model**.

The Random Intercept Model

We have that

$$Y_i = \beta_0 + b_{0i} + \tilde{X}_i \beta + \epsilon_i,$$

with $b_{0i} \sim N(0, \sigma_b^2)$ and $\epsilon_i \sim N(0, G_i = \sigma^2 I)$.

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This gives

$$\text{cov}(Y_{ij}, Y_{il}) = \frac{\sigma_b^2}{\sigma_b^2 + \sigma^2}.$$

As a result, a random intercept model imposes the compound symmetry assumption!

The Random Intercept and Slope Model

If instead of *just* a random intercept, we also include a **random time slope** we get

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Here, D will be given by the variance of each b_{0i} and b_{1i} , as well as by the covariance between these terms.

The within-subject correlation will be time dependent in this model automatically!

Parameter Estimation and Hypothesis Testing

This is a Parametric Model

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This will give the (familiar) asymptotic results where

$$\hat{\beta} \sim N \left(\beta, \left[\sum_{i=1}^n X_i' V_i^{-1}(\theta) X_i \right]^{-1} \right),$$

with $V_i(\theta) = \text{var}(Y_i) = Z_i D Z_i' + G_i$.

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- ▶ Can use the standard **information criteria** as well!

Response Prediction (BLUPs)

Estimation versus Prediction

We saw **estimation** of the parameters β , but the b_i are random!
As a result, we must **predict** them.

The best[†] predictor for b_i will be $E[b_i|Y_i]$, a quantity that we call the **best linear unbiased prediction** (or BLUP).

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- ▶ Once estimated, we can estimate outcomes as

$$\hat{Y}_i = X_i\hat{\beta} + Z_i\hat{b}_i = \dots = \hat{G}_i\hat{V}_i^{-1}X_i\beta + [I - \hat{G}_iV_i^{-1}]Y_i.$$

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- ▶ This is a **weighted average** between the estimated population mean ($X_i\hat{\beta}$) and the individual observation Y_i .
- ▶ When G_i is **large** (more within-subject variation than between) there is more weight to the population average, and vice-versa.

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- ▶ Two basic, common models (**random intercept** and **random intercept and slope**) capture correlation structures that we have previously seen.
- ▶ We can use the **BLUP** to estimate individual effects, as-is necessary.

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- ▶ They make distributional assumptions (GEEs did not).
- ▶ They may imply overly complex structures at the marginal level.